Articles

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Non-Communicable Diseases Unit, Department of Public Health, Institute of Tropical Medicine, Antwerp, Belgium (Prof LL Peñalvo PhD. D Sagastume MSc, E Mertens PhD): Friedman School of Nutrition Science and Policy, Tufts University, Boston, MA, USA (Prof J L Peñalvo, J Smith PhD, E Bishop MSc, I Onopa MSc, P Shi MSc. Prof R Micha PhD. Prof D Mozaffarian PhD); Department of Health and Nutritional Sciences. Institute of Technology Sligo, Sligo, Ireland (I Uzhova PhD); Bell Institute of Health and Nutrition, General Mills, Minneapolis, MN, USA (I Smith): George Institute for Global Health, Faculty of Medicine, University of New South Wales, Sydney, NSW. Australia (Prof I H Y Wu PhD): Department of Food Science and Human Nutrition, University of Thessalv. Thessaly, Greece (Prof R Micha)

Correspondence to: Prof José L Peñalvo, Non-Communicable Diseases Unit, Department of Public Health, Institute of Tropical Medicine, 2000 Antwerp, Belgium jpenalvo@itg.be

habits, overweight, and cardiometabolic health: a systematic review and meta-analysis

Effectiveness of workplace wellness programmes for dietary

José L Peñalvo, Diana Sagastume, Elly Mertens, Irina Uzhova, Jessica Smith, Jason H Y Wu, Eve Bishop, Jennifer Onopa, Peilin Shi, Renata Micha, Dariush Mozaffarian

Summary

Background The workplace offers a unique opportunity for effective health promotion. We aimed to comprehensively study the effectiveness of multicomponent worksite wellness programmes for improving diet and cardiometabolic risk factors.

Methods We did a systematic literature review and meta-analysis, following PRISMA guidelines. We searched PubMed-MEDLINE, Embase, the Cochrane Library, Web of Science, and Education Resources Information Center, from Jan 1, 1990, to June 30, 2020, for studies with controlled evaluation designs that assessed multicomponent workplace wellness programmes. Investigators independently appraised the evidence and extracted the data. Outcomes were dietary factors, anthropometric measures, and cardiometabolic risk factors. Pooled effects were calculated by inverse-variance random-effects meta-analysis. Potential sources of heterogeneity and study biases were evaluated.

Findings From 10169 abstracts reviewed, 121 studies (82 [68%] randomised controlled trials and 39 [32%] quasiexperimental interventions) met the eligibility criteria. Most studies were done in North America (57 [47%]), and Europe, Australia, or New Zealand (36 [30%]). The median number of participants was 413.0 (IQR 124.0-904.0), and median duration of intervention was 9.0 months (4.5-18.0). Workplace wellness programmes improved fruit and vegetable consumption (0.27 servings per day [95% CI 0.16 to 0.37]), fruit consumption (0.20 servings per day [0.11 to 0.28]), body-mass index (-0.22 kg/m² [-0.28 to -0.17]), waist circumference (-1.47 cm [-1.96 to -0.98]), systolic blood pressure (-2.03 mm Hg [-3.16 to -0.89]), and LDL cholesterol (-5.18 mg/dL [-7.83 to -2.53]), and to a lesser extent improved total fat intake (-1.18% of daily energy intake [-1.78 to -0.58]), saturated fat intake (-0.70% of daily energy [-1.22 to -0.18]), bodyweight (-0.92 kg [-1.11 to -0.72]), diastolic blood pressure (-1.11 mm Hg [-1.78 to -0.44]), fasting blood glucose (-1.81 mg/dL [-3.33 to -0.28]), HDL cholesterol (1.11 mg/dL [0.48 to 1.74]), and triglycerides (-5.38 mg/dL [-9.18 to -1.59]). No significant benefits were observed for intake of vegetables (0.03 servings per day [95% CI -0.04 to 0.10]), fibre (0.26 g per day [-0.15 to 0.67]), polyunsaturated fat (-0.23% of daily energy [-0.59 to 0.13]), or for body fat (-0.80% [-1.80 to 0.21]), waist-to-hip ratio (-0.00 ratio [-0.01 to 0.00]), or lean mass (1.01 kg [-0.82 to 2.83]). Heterogeneity values ranged from 46.9% to 91.5%. Betweenstudy differences in outcomes were not significantly explained by study design, location, population, or similar factors in heterogeneity analyses.

Interpretation Workplace wellness programmes are associated with improvements in specific dietary, anthropometric, and cardiometabolic risk indicators. The heterogeneity identified in study designs and results should be considered when using these programmes as strategies to improve cardiometabolic health.

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Introduction

Adults spend most of their weekday waking hours at work, and thus the workplace offers a unique opportunity to promote health. Both the World Health Assembly in 2004' and the UN high-level meeting on the prevention and control of non-communicable diseases in 2011 called on the private sector to promote enabling environments and worksite wellness programmes for healthy behaviours among workers.² In addition to improved employee health, workplace wellness programmes might benefit companies through higher employee satisfaction, increased loyalty, improved productivity, and lower health-care costs.³⁴ In 2017, WHO identified workplace wellness programmes as a best-buy option for the prevention and control of non-communicable diseases including mental health.⁵

Although narrative reviews have suggested benefits of workplace wellness programmes for lifestyle behaviours and cardiometabolic health,⁶⁻¹⁰ few quantitative metaanalyses have been done to identify magnitudes of

Research in context

Evidence before this study

We did a preliminary review of scientific and policy statements and quidance from organisations such as WHO, the Institute of Medicine, US Department of Health and Human Services, and other similar international and national health agencies using synonyms and combinations of keywords for workplaces, health promotion, and health outcomes. We found extensive but dispersed evidence suggesting potential health benefits of workplace wellness programmes in improving several lifestyle and cardiometabolic risk factors among staff members. Mostly narrative reviews, and only a few quantitative meta-analyses, have been done to identify magnitudes of benefits, but these were limited to specific populations or targets and outcomes of workplace wellness programmes. Therefore, we searched PubMed-MEDLINE, Embase, Cochrane Library, Web of Science, and Education Resources Information Center, from Jan 1, 1990, to June 30, 2020, for studies with controlled evaluation designs published in English that assessed multicomponent workplace wellness programmes. We aimed to provide the most comprehensive review of workplace wellness programme designs, quantify their effectiveness in improving several risk factors and outcomes, and identify the major drivers of impact.

Added value of this study

We systematically extracted, reviewed, and catalogued information on the design and characteristics of available workplace wellness programmes that have been assessed for

benefits for specific outcomes or employee-related or intervention-related factors that may influence effectiveness.¹¹⁻¹³ Also, most of these analyses were limited to a specific population, intervention targets (eg, diet, physical activity), or outcomes (eg, fruit intake, bodyweight, glycaemia), with few included studies (eg, ten total studies11) in each. To our knowledge, no previous systematic review and meta-analysis has comprehensively assessed the effect of workplace wellness programmes on a broad set of dietary and dietrelated health indicators. Additionally, few studies have quantitatively explored potential heterogeneity in outcomes, for instance based on employee characteristics or intervention components.14 To address these gaps in knowledge, we did a systematic review and meta-analysis of the effects of multicomponent workplace wellness programmes on dietary behaviours and major dietrelated cardiometabolic risk factors.

Methods

Search strategy and selection criteria

This systematic review and meta-analysis was done in accordance with the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) guidelines.¹⁵ The protocol is presented in the appendix (pp 5–6).

effectiveness in studies with controlled designs. From the 121 studies reviewed, we identified enough information to calculate pooled estimates for 20 different outcomes, of which 13 were found to be positively affected by workplace wellness programmes, especially fruit and vegetable consumption and markers of adiposity. This research compiles and provides novel findings that integrate the scientific evidence accumulated during the past 30 years on the effects of workplace wellness programmes on dietary factors, body anthropometrics, and cardiometabolic risk.

Implications of all the available evidence

International health organisations recommend workplace programmes as one of the key strategies for improving population health. Our work supports this recommendation by demonstrating that workplace programmes are associated with improvements in fruit and vegetable consumption, bodyweight, and body-mass index, and potentially improvements in cardiometabolic risk factors. Promoting a healthy diet and increasing physical activity among workers, by means of education strategies and accompanying improvements in the food environment of the workplace, were the most common elements in these interventions. These results support the implementation of workplace programmes, and highlight the need for further research to identify the most cost-effective approaches, and how to tailor these programmes in different socioeconomic contexts.

Reports from health agencies related to workplace wellness programmes were first reviewed by the investigators for development of the protocol, contextualisation of the research, and identification of key scientific literature. This preliminary review included policy statements and guidance from the Institute of Medicine, WHO, Centers for Disease Control and Prevention, US Department of Health and Human Services, and other similar international, national, and local agencies.

The systematic search was done in PubMed-MEDLINE, Embase, the Cochrane Library, Web of Science, and the Education Resources Information Center, for studies published in English from Jan 1, 1990, to June 30, 2020. The search strategy was developed and implemented under the guidance of experts of library services from Tufts University (Boston, MA, USA). The search terms included different synonyms and combinations of words for workplaces, health promotion, weight loss, diet, and cardiometabolic factors (appendix pp 7-8). Online searches were supplemented by hand searches of reference lists of the first 20 related articles suggested in PubMed for each of the final included articles. Titles and abstracts were screened in duplicate and, for all potentially relevant articles, full-text manuscripts were retrieved for further review and eligibility check.

Studies were selected if they were interventional controlled trials, either randomised or quasiexperimental with external control or comparison group. Commentaries, protocols, or review articles were also assessed as sources of potential references. Studies were excluded if they were quasi-experimental without an external control or comparison group, observational, ecological, theoretical (laboratory experiments), or simulation (modelling) study designs; or duplicate publications from the same study (after fulltext review). Studies were eligible if they reported any worksite-based intervention that targeted the overall workplace. Studies of non-employed individuals, children, or an intervention aimed solely at disease management (eg, control of existing type 2 diabetes) were excluded. The study setting could be any type of workplace, apart from schools targeting students and non-workplace organisations (eg, community centres or religious organisations).

Eligible studies assessed multicomponent workplace wellness programmes based on two or more intervention components that targeted improved health, such as the use of educational messages, cafeteria or vending machine interventions, promotion of stair use, financial incentives, changes to health insurance policies, or improved accessibility or discounts for gym memberships. Only multicomponent programmes were included because there is evidence of this being the most effective approach, particularly when targeting diet and physical activity,^{16,17} and multicomponent workplace wellness programmes are recommended by WHO for the prevention of non-communicable diseases.18,19 Studies were excluded if they assessed single-component programmes or tailored individual-level interventions not part of a multicomponent intervention, routine workplace health screening programmes without additional multicomponent intervention, work-life balance programmes that did not target improvements in health (eg, stress management), or smoking cessationonly programmes.

We included studies if they reported any of the following outcomes: change in dietary habits (measured by food frequency questionnaires, 24-hour recall, or dietary records), markers of adiposity (eg, bodyweight, body-mass index [BMI], waist circumference, skinfolds, body fat percentage), cardiometabolic risk factors (eg, blood pressure, lipids, glucose, insulin), cardiovascular risk scores (eg, Framingham risk score), or disease outcomes (eg, diabetes) if available. Studies had to report the differential change (intervention vs control) in the outcome of interest plus a measure of uncertainty for the reported difference, or information to compute these numbers. We did not include studies if they only reported changes in knowledge or attitudes regarding health, diet, or physical activity; changes in health-care costs to the company; absenteeism; or changes in mental health, general wellbeing, or quality of life.

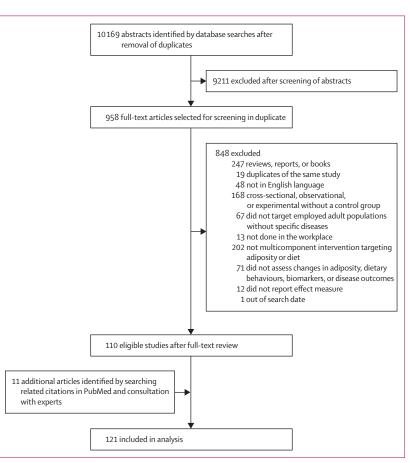


Figure 1: Study selection

Data extraction

All of the following steps were done in duplicate by two investigators. Studies selected for full-text review were evaluated independently for inclusion, with any differences resolved by consensus. Data were extracted individually using standardised electronic templates. The extracted information is detailed in the appendix (p 8). Study quality was assessed using a previously established and utilised scoring system based on five criteria: study design, assessment of exposure, assessment of outcome, control for confounding, and evidence of selection bias (appendix p 9).²⁰⁻²² Each criterion received a score of 0 or 1 (with 1 being better) and an overall score was calculated as the sum of individual scores; with 0-3 considered as lower quality and 4-5 considered higher quality. Balance between intervention and control groups, differential attrition rate, and fit of the statistical model chosen to the study design, in terms of unit of analysis, covariate adjustment, and uncertainty of the estimate were also used to assess study quality. To address and assess the variability in the intervention components and targets used across the studies and facilitate identification of potentially more influential intervention design characteristics of workplace wellness

programmes, individual components and targets of the intervention (eg, focus on physical activity or dietary habits) were classified into different domains. These categories (appendix p 9) were assessed in exploration of heterogeneity. Missing data or unclear information were attempted to be resolved by direct contact with authors; when feedback was not received, assumptions were discussed and agreed upon by the investigators (appendix pp 10–11). Differences in data extraction and bias assessment between investigators were infrequent

| | All studies (n=121) | Randomised controlled trials (n=82) | Quasi-experimental studies (n=39) |
|--|---------------------|---|--------------------------------------|
| Study details | | | |
| Publication date | | | |
| Before 2000 | 18 (15%) | 13 (16%) | 5 (13%) |
| 2000-09 | 31 (26%) | 14 (17%) | 17 (44%) |
| 2010 or later | 72 (60%) | 55 (67%) | 17 (44%) |
| Country | | | |
| North America (US and Canada) | 57 (47%) | 46 (56%) | 11 (28%) |
| Europe*, Australia, New Zealand | 36 (30%) | 22 (27%) | 14 (36%) |
| Asia† | 22 (18%) | 9 (11%) | 13 (33%) |
| Other‡ | 6 (5%) | 5 (6%) | 1 (3%) |
| Unit of randomisation | | | |
| Individual | | 24 (29%) | |
| Cluster | | 58 (71%) | |
| Lost to follow-up (%) | 26.0% (20.9) | 24.4% (20.2) | 29.6% (22.2) |
| Median (IQR) | 20.9% (12.0-37.4) | 19.8% (12.0–34.6) | 27.9% (12.2-43.3) |
| Bias score§ | 3.4 (1.0) | 3.8 (0.7) | 2.4 (0.9) |
| Median (IQR) | 4.0 (3.0-4.0) | 4.0 (3.0-4.0) | 2.0 (2.0–3.0) |
| Workplace settings | | | |
| Type¶ | | | |
| Office | 19 (16%) | 13 (16%) | 6 (16%) |
| Hospital | 15 (13%) | 14 (18%) | 1 (3%) |
| School | 5 (4%) | 5 (6%) | 0 |
| Factory | 22 (19%) | 11 (14%) | 11 (30%) |
| Mixed or other | 56 (48%) | 37 (46%) | 19 (51%) |
| Number of sites | 9.4 (14.2) | 10.5 (15.3) | 7.2 (11.2) |
| Median (IQR) | 4.0 (2.0–13.0) | 4.0 (2.0–16.0) | 3.0 (2.0–7.5) |
| Employee characteristics | | | |
| Number of participating employees | 1231 (3147) | 886 (1532) | 1964 (4870) |
| Median (IQR) | 413.0 (124.0–904.0) | 314.0 (98.0–782.0) | 547.0 (253.0-1749.0) |
| Age, years | 42.7 (5.7) | 43.0 (5.8) | 41.8 (5.4) |
| Median (IQR) | 43.0 (38.9-46.0) | 43.6 (39.9–46.5) | 42.0 (37.3-44.6) |
| Sex, percentage male** | 51.4% (30.8) | 48.4% (31.4) | 58.0% (29.1) |
| Median (IQR) | 49·2% (24·6–76·1) | 46.8% (22.6–74.4) | 60.1% (37.7-77.1) |
| Race or ethnicity, percentage white†† | 56.0% (34.2) | 62.9% (30.8) | 36.3% (36.7) |
| Median (IQR) | 61.3% (32.4-89.8) | 73.1% (42.4–90.8) | 38.1% (0.0-68.3) |
| BMI (kg/m²) | 28.0 (3.3) | 28.7 (3.5) | 26.5 (2.6) |
| Median (IQR) | 27.5 (25.5–29.7) | 28.2 (25.7–31.6) | 26.4 (25.2–28.1) |

and resolved by the duplicate abstractors and a third investigator.

For randomised and quasi-experimental interventional studies, the differential (between-group, intervention vs comparator) change at follow-up adjusted for baseline values and relevant covariates was extracted when available or calculated based on the information reported. Statistical uncertainty (standard error, SE) of effect sizes was extracted or calculated based on other statistics, or in the absence of other information, assumed based on information provided in the article (appendix pp 12–15). For paired observations without reported covariance, we used a correlation coefficient of 0.9 if no loss to followup, 0.5 if loss to follow-up, and 0 if independent samples. Effects were extracted as continuous effect sizes when available, and as other effect sizes otherwise (eg, percentage meeting a cutpoint, odds ratio), together with their statistical uncertainty. Multiple trial groups were handled as separate intervention groups from the same study and were included as separate estimates in the meta-analyses; subgroup findings from the same intervention group (eg, by sex, age) were first combined using study-specific meta-analysis. Effect sizes were standardised to servings per day for fruits and vegetables, g per day for dietary fibre, percentage daily energy for fats and fatty acids, kg for bodyweight and lean mass, kg/m² for BMI, cm for waist circumference, percentage for body mass, mm Hg for blood pressure, and mg/dL for plasma glucose and lipid fractions. Information on outcomes related to intake of sodium, monounsaturated fatty acids, red meat, and whole grains; insulin and apolipoprotein concentrations; disease risk scores (eg, Framingham risk score); and disease outcomes (eg, cardiovascular disease) was scarce, and not enough for meta-analysis; thus, these outcomes were only included in the qualitative assessment of the evidence.

Data analysis

Study-specific effect sizes and corresponding SE were pooled using inverse-variance random-effects metaanalysis to estimate an overall summary effect size (95% CI) and to obtain the corresponding forest plots for each of the outcomes of interest. All analyses were done with Stata (release 16/SE).

Cochran's Q and I^2 statistics were used to assess between-study heterogeneity.²³ For outcomes with ten or more study estimates, univariate and multivariate metaregressions explored prespecified sources of potential heterogeneity including study publication, location, study design, bias score, type of workplace, duration of the intervention, number of intervention components, mean employee age, and sex. Because analyses of heterogeneity were exploratory across multiple comparisons, a Bonferroni-adjusted p<0.001 was considered to be significant in multivariate meta-regressions. Potential small-study effects or public bias were assessed by visual inspection of funnel plots, and computation of Egger's test and Begg's test.²⁴ The trim-and-fill method²⁵ was used to adjust for potential small-study effects or publication bias and compare with unadjusted findings. To further evaluate the consistency of the results depending on study design, a sensitivity analysis was done excluding quasi-experimental studies (ie, including randomised controlled trials only) for both meta-analyses of effects, and trim-and-fill for small-study effects.

Role of the funding source

The funder of the study had no role in study design, data collection, data analysis, data interpretation, or writing of the report.

Results

Of 10169 identified abstracts and 958 articles identified for full-text review, 121 interventional studies met the inclusion and exclusion criteria, including 82 (68%) randomised controlled trials and 39 (32%) quasiexperimental trials (figure 1). Study characteristics are summarised in table 1, with details of each individual study reported in the appendix (pp 16-31). Most studies (72 [60%] of 121) were published from 2010 onwards. Studies were from North America (57 [47%]), western Europe, Australia, and New Zealand (36 [30%]), Asia (22 [18%]), and other countries (Brazil, South Africa, and Tunisia; 6 [5%]). Workplace settings included factories, offices, hospitals, and schools (employees), and mixed settings were predominant (56 [48%] of 117). Numbers of participating employees varied (median 413.0 employees [IQR 124.0-904.0]), as did duration of the intervention (median 9.0 months $[4 \cdot 5 - 18 \cdot 0]$). Behavioural intervention targets included dietary habits (94 [78%] of 121), physical activity (81 [67%]), and weight loss (53 [44%]), with most studies (101 [84%]) having more than one target. More than 50 individual intervention components were identified across the retrieved studies, which were classified into nine intervention domains (appendix p 9): screening, individual education, group education, food environment, labelling, financial incentives, physical activity, self-awareness, and other (eg, employees' advisory committees). Identified trials compared workplace wellness programme interventions with a comparison group that was either a base case (usual care) control or a less intensive intervention (eg, basic education about healthy diet only). Assessment of bias scores were higher (more favourable) for randomised controlled interventions (mean 3.8 [SD 0.7]) versus nonrandomised designs $(2 \cdot 4 [0 \cdot 9])$.

Pooled quantitative estimates were derived for seven dietary factors, six anthropometric measures, and seven clinical risk factors (table 2). The most assessed outcome among anthropometric factors was BMI (67 intervention groups); for dietary habits, total fruit and vegetable consumption as well as total fat were the most assessed outcomes (both with 18 intervention groups);

| | All studies (n=121) | Randomised controlled trials (n=82) | Quasi-experimental studies (n=39) |
|----------------------------|---------------------|---|--------------------------------------|
| (Continued from previous p | bage) | | |
| Intervention characteristi | cs | | |
| Duration, months | 13·3 (13·8) | 10.1 (8.6) | 20.0 (19.3) |
| Median (IQR) | 9.0 (4.5–18.0) | 6.0 (3.0–15.0) | 12.0 (8.0–24.0) |
| Primary specified target‡‡ | | | |
| Diet | 94 (78%) | 60 (73%) | 34 (87%) |
| Physical activity | 81 (67%) | 51 (62%) | 30 (77%) |
| Weight management | 53 (44%) | 36 (44%) | 17 (44%) |
| Other | 76 (63%) | 52 (63%) | 24 (62%) |
| Number of components | | | |
| 2 | 12 (10%) | 6 (7%) | 6 (15%) |
| 3 | 31 (26%) | 16 (20%) | 15 (38%) |
| 4 | 22 (18%) | 15 (18%) | 7 (18%) |
| 5 | 14 (12%) | 11 (13%) | 3 (8%) |
| 6 | 10 (8%) | 9 (11%) | 1(3%) |
| 7 | 5 (4%) | 5 (6%) | 0 |
| 8 | 8 (7%) | 6 (7%) | 2 (5%) |
| 9 | 9 (7%) | 6 (7%) | 3 (8%) |
| 10 | 10 (8%) | 8 (10%) | 2 (5%) |

Data are n (%) or mean (SD), unless otherwise specified. --=not applicable. BMI=body-mass index. *Europe includes the Netherlands, Norway, Denmark, Belgium, the UK, Finland, Germany, Iceland, Ireland, and Switzerland. †Asia includes Japan, Taiwan, South Korea, Malaysia, Singapore, India, China, Jordan, and Bangladesh. ‡Other countries are Brazil, Tunisia, and South Africa. §The bias score assigned to each study ranged from 1 to 5. This value represents the mean (SD) of the score and the corresponding median (IQR). ¶Denominator is 117; four studies did not report the type of workplace. Workplace settings were classified into the following types: office, hospital, school, factory, mixed, or others. Mixed is defined as a combination of the prespecified workplaces, whereas others differ from the prespecified categories. ||A mean of 1188 participants (SD 1739) were cluster randomised and 171 (183) participants were individually randomised. **Data on number of women were not extracted and therefore are not available. ††When available, data on race or ethnicity usually referred to white people, and so this group was chosen to represent this variable in the meta-analyses. ##Primary specified target indicates the most frequent, but not limited to, intervention focus per outcome identified subjectively by the investigators (diet quality, weight loss, physical activity, and other such as reduction in cardiovascular disease risk factors, smoking cessation, stress reduction, diabetes, or cancer prevention); for instance the primary specified target for the interventions reporting fruit intake was diet quality (93%), meaning diet quality was the most frequent target, but not exclusively because the intervention could also target physical activity or weight loss less frequently.

Table 1: Main characteristics of the 121 studies reporting assessment of the effect of workplace wellness programmes

and for clinical risk factors, blood pressure was the most assessed outcome (41 intervention groups).

Among dietary factors, workplace wellness programmes increased intake of total fruits and vegetables (0·27 servings per day [95% CI 0·16 to 0·37]) and intake of fruits (0·20 servings per day [0·11 to 0·28]; figure 2), and decreased intake of total fat ($-1\cdot18\%$ of daily energy intake [$-1\cdot78$ to $-0\cdot58$]) and saturated fat (-0.70% of daily energy [$-1\cdot22$ to $-0\cdot18$]). No significant changes were identified for intake of vegetables (0·03 servings per day [95% CI -0.04 to 0·10]), dietary fibre (0·26 g per day [-0.15 to 0·67]), or polyunsaturated fat (-0.23% of daily energy intake [-0.59 to 0·13]). Workplace wellness programmes significantly reduced BMI (-0.22 kg/m² [95% CI -0.28 to -0.17]), bodyweight (-0.92 kg [-1.11 to -0.72]), and waist circumference (-1.47 cm [-1.96 to -0.98]; figure 3). A numerical decrease was seen for body fat

| | Number of studies (number of intervention groups)* | Number of intervention groups from RCT (%) | Number of participants, median (IQR) | Primary specified targets, n (%)† | Duration, months (SD) | Pooled effect size (95% Cl)‡ | ², % | p asymmetry (Egger's test) |
|---|--|---|--|---|--------------------------|---------------------------------|------|-------------------------------|
| Dietary habits | | | | | | | | |
| Fruits and vegetables, servings per day | 16 (18) | 15 (83%) | 550 (397–1359) | 18 (100%)§ | 15.6 (8.9) | 0·27 (0·16 to 0·37) | 88.7 | 0.0004 |
| Fruits, servings per day | 13 (15) | 13 (87%) | 430 (257–730) | 14 (93%)§ | 6.5 (5.4) | 0·20 (0·11 to 0·28) | 59·9 | 0.0092 |
| Vegetables, servings per day | 12 (14) | 10 (71%) | 419 (314–515) | 13 (93%)§ | 6.8 (5.4) | 0.03 (-0.04 to 0.10) | 75.6 | 0.010 |
| Fibre, g per day | 8 (15) | 11 (73%) | 433 (174–850) | 15 (100%)§ | 11.8 (6.6) | 0·26 (-0·15 to 0·67) | 61.3 | 0.63 |
| Total fat, % daily energy | 14 (18) | 12 (67%) | 478 (362-850) | 18 (100%)§ | 11.7 (7.4) | –1·18 (–1·78 to –0·58) | 80.1 | 0.44 |
| Saturated fat, % energy | 4 (6) | 2 (33%) | 850 (770–850) | 6 (100%)§ | 11.0 (7.4) | -0.70 (-1.22 to -0.18) | 46.9 | 0.046 |
| Polyunsaturated fat, % energy | 3 (3) | 2 (67%) | 770 (186–3076) | 3 (100%)§ | 14.0 (10.5) | -0·23 (-0·59 to 0·13) | 62.6 | 0.23 |
| Anthropometric measures | | | | | | | | |
| BMI, kg/m² | 57 (67) | 41 (61%) | 447 (110–904) | 53 (79%)¶ | 14.0 (12.1) | -0.22 (-0.28 to -0.17) | 86.9 | 0.0013 |
| Bodyweight, kg | 47 (59) | 35 (59%) | 269 (77-850) | 45 (76%) | 10.0 (8.7) | -0·92 (-1·11 to -0·72) | 86.8 | 0.0068 |
| Waist circumference, cm | 31 (37) | 18 (49%) | 265 (95–553) | 30 (81%) | 10.3 (12.6) | -1·47 (-1·96 to -0·98) | 81.9 | 0.061 |
| Body fat, % | 11 (13) | 10 (77%) | 60 (58–77) | 13 (100%) | 4.8 (2.9) | -0.80 (-1.80 to 0.21) | 87.0 | 0.84 |
| Waist-to-hip, ratio | 6 (8) | 8 (100%) | 169 (45–592) | 7 (88%)** | 10.4 (12.6) | -0.00 (-0.01 to 0.00) | 79.8 | 0.17 |
| Lean mass, kg | 4 (4) | 4 (100%) | 50 (30–189) | 4 (100%) | 3·3 (0·5) | 1.01 (-0.82 to 2.83) | 89.8 | 0.45 |
| Cardiometabolic risk factors | | | | | | | | |
| Systolic blood pressure, mm Hg | 34 (41) | 22 (54%) | 228 (102-817) | 33 (80%)¶ | 13.4 (18.0) | -2.03 (-3.16 to -0.89) | 89.5 | 0.019 |
| Diastolic blood pressure, mm Hg | 34 (41) | 22 (54%) | 228 (102–817) | 33 (80)¶ | 13.4 (18.0) | –1·11 (–1·78 to –0·44) | 74·7 | 0.51 |
| Fasting glucose, mg/dL | 21 (26) | 12 (46%) | 190 (70–1371) | 23 (88%) | 12.7 (11.6) | -1.81 (-3.33 to -0.28) | 91·5 | 0.22 |
| HDL cholesterol, mg/dL | 29 (32) | 15 (47%) | 190 (74–532) | 28 (88%) | 14.2 (16.0) | 1·11 (0·48 to 1·74) | 89.6 | 0.077 |
| LDL cholesterol, mg/dL | 20 (22) | 11 (50%) | 162 (70–447) | 21 (95%) | 11·9 (13·2) | -5·18 (-7·83 to -2·53) | 87.5 | 0.0022 |
| Triglycerides, mg/dL | 23 (26) | 12 (46%) | 142 (61–490) | 23 (88%) | 12·3 (14·6) | -5·38 (-9·18 to -1·59) | 67.8 | 0.013 |
| Total cholesterol, mg/dL | 32 (36) | 17 (47%) | 292 (104–1138) | 31 (86%)¶ | 15.7 (15.2) | –1·75 (–2·59 to –0·91) | 83.6 | 0.0056 |

Table 2: Pooled estimates of the effect (change) of workplace wellness programmes on dietary habits, anthropometric measurements, and clinical parameters

percentage (-0.80% [95% CI -1.80 to 0.21]) and an increase for lean mass (1.01 kg [-0.82 to 2.83]), but these changes were not significant. No significant changes were seen in waist-to-hip ratio (-0.00 ratio [95% CI -0.01 to 0.00]); eight intervention groups).

Significant improvements were identified in all cardiometabolic risk factors (table 2). Systolic blood pressure declined by $2 \cdot 03 \text{ mm Hg}$ (95% CI – $3 \cdot 16 \text{ to} -0 \cdot 89$) and diastolic blood pressure by $1 \cdot 11 \text{ mm Hg}$ (– $1 \cdot 78 \text{ to} -0 \cdot 44$). Fasting glucose decreased by $1 \cdot 81 \text{ mg}$ / dL (95% CI – $3 \cdot 33 \text{ to} -0 \cdot 28$), LDL cholesterol by $5 \cdot 18 \text{ mg/dL}$ (– $7 \cdot 83 \text{ to} -2 \cdot 53$; figure 4), triglycerides by $5 \cdot 38 \text{ mg/dL}$ (– $9 \cdot 18 \text{ to} -1 \cdot 59$), and HDL cholesterol increased by $1 \cdot 11 \text{ mg/dL}$ ($0 \cdot 48 \text{ to} 1 \cdot 74$). Forest plots for all outcomes are presented in the appendix (pp 32–41).

Substantial levels of heterogeneity (I^2 >60%) were observed for most of the outcomes, except for fruit intake and saturated fat intake, which showed moderate heterogeneity (I2<60%). Results for prespecified sources of heterogeneity are presented in the appendix (pp 42–57). The multivariate meta-regressions corrected for multiple comparisons suggested that changes towards healthier food environments seem to be a relevant intervention component for improvements in blood lipids, particularly in LDL cholesterol (p_{interaction}<0.001). Stratified analysis for this finding could not be reliably done because of the simultaneity of several intervention components (eg, an intervention includes, concurrently, a food environment component such as providing free fruit at work, and a physical activity component such as using a pedometer). No other statistically significant effect modifier was identified by the exploratory analyses, indicating that the heterogeneity in several outcomes could not be explained by the potential moderating factors that we investigated here. Scores for assessment of bias were not observed to be a significant source of

| | Country | Number of participants | Randomised | Intervention domains | Duration (months) | | Effect size, servings per day (95% CI) | Weight (%) |
|---------------------------------|-------------|---------------------------|------------|-------------------------|----------------------|------------|--|---------------|
| Wierenga (2014) ⁵¹ | Netherlands | 2139 | No | ACDEGHI | 12 | | -0.04 (-0.29 to 0.21) | 6.58 |
| van Berkel (2014) ⁴⁹ | Netherlands | 257 | Yes | BCDGH | 6 | + | -0.02 (-0.24 to 0.20) | 7.32 |
| Wilson (2016)43 | USA | 362 | Yes | BCDE | 6 | - | 0.06 (-0.01 to 0.13) | 13.38 |
| Hunt (1993) ⁵⁰ | USA | 2365 | Yes | ABCEI | 15 | | 0·13 (-0·07 to 0·33) | 8.11 |
| Viester (2018)46 | Netherlands | 314 | Yes | ABG | 6 | | 0·17 (-0·04 to 0·38) | 7.79 |
| Hutchinson (2013) ⁴⁵ | Australia | 55 | Yes | CD | 1 | | 0·17 (-0·34 to 0·69) | 2.37 |
| Steenhuis (2004) ^{44*} | Netherlands | 508 | Yes | BEI | 1 | | 0·20 (0·07 to 0·33) | 10.84 |
| Steenhuis (2004) ^{44*} | Netherlands | 430 | Yes | BEI | 1 | | 0.20 (0.06 to 0.34) | 10.32 |
| Steenhuis (2004) ^{44*} | Netherlands | 505 | Yes | BDEI | 1 | - | 0·20 (0·07 to 0·33) | 10.78 |
| Strijk (2012)47 | Netherlands | 730 | Yes | BDGI | 6 | | 0·39 (0·10 to 0·68) | 5.50 |
| Morgan (2011) ⁴⁸ | Australia | 86 | Yes | ABCFGH | 3 | | 0.40 (0.00 to 0.80) | 3.54 |
| Campbell (2002) ³⁹ | USA | 859 | Yes | ABCDG | 18 | | 0·40 (0·11 to 0·69) | 5.55 |
| Lassen (2011) ³⁶ | Denmark | 206 | Yes | BDEI | 6 | | 0.42 (-0.30 to 1.14) | 1.30 |
| Kuehl (2014) ³⁰ | USA | 408 | Yes | AC | 3 | | 0.50 (0.25 to 0.75) | 6.51 |
| Engbers (2006)42 | Netherlands | 515 | No | EI | 12 | • | 3·57 (1·03 to 6·11) | 0.12 |
| Overall (1²=59·9%, p=0·0 | 0014) | | | | | \diamond | 0·20 (0·11 to 0·28) | 100.00 |

Figure 2: Forest plot of intake of fruits

Intervention domains correspond to screening (A), individual education (B), group education (C), food environment (D), labelling (E), financial incentives (F), physical activity (G), self-awareness (H), and others (I). Weights are from random-effects analysis. *Different intervention groups from the same study (Steenhuis 2004).

differential effects (appendix pp 42–44), nor a source of heterogeneity.

Visual inspection of funnel plots and Egger's regression tests was done to assess small-study effects or publication bias (appendix pp 58-67). Neither plot asymmetry nor significant Egger's tests were observed for several outcomes, including intakes of dietary fibre, total fat, or polyunsaturated fat, waist circumference, body fat percentage, waist-to-hip ratio, lean mass, diastolic blood pressure, fasting plasma glucose, or HDL cholesterol. Egger's test suggested potential small-study effects or publication bias for intakes of fruits and vegetables (p=0.0004), fruits (p=0.0092), or vegetables (p=0.010), and saturated fat (p=0.046), BMI (p=0.0013), bodyweight (p=0.0068), systolic blood pressure (p=0.019), LDL cholesterol (p=0.0022), triglycerides (p=0.013), and total cholesterol (p=0.0056); further supported by visual asymmetry in the funnel plots. The trim-and-fill method was used to adjust for these potential effects and compared with the unadjusted pooled effect sizes (appendix p 68). After adjustment, the estimates of the effect of workplace wellness programmes on dietary intake of total fruits and vegetables (0.05 servings per day [95% CI -0.05 to 0.16]) and saturated fat (-0.31% of daily energy intake [-0.87 to 0.25]) were no longer statistically significant; intake of vegetables remained not significant (-0.06 servings per day [-0.13 to 0.02]) and the effect on intake of fruits (0.12 servings per day [0.11 to 0.28]) was smaller but still significant. Similarly, effect sizes were smaller but still significant for BMI (-0.12 kg/m² [95% CI -0.18 to -0.06]) and bodyweight (-0.52 kg [-0.72 to -0.31]). Effects of workplace wellness programmes were no longer significant for systolic blood pressure (-0.06 mm Hg [95% CI -1.31 to 1.20]), LDL cholesterol (-0.41 mg/dL [-3.00 to 2.18]), triglycerides (-2.14 mg/dL [-6.39 to 2.11]), and total cholesterol (-0.30 mg/dL [-1.20 to 0.59]) after correction.

In sensitivity analysis considering randomised controlled trials only (appendix p 69), overall similar or slightly higher pooled effects were observed for dietary and anthropometric outcomes (eg, intake of fruits and vegetables 0.34 servings per day [95% CI 0.21 to 0.47]). For cardiometabolic risk factors, however, estimates were slightly smaller or became non-significant for fasting glucose, HDL cholesterol, and triglycerides. Small-study effect or publication bias was suggested for the estimates of fruits and vegetables intake combined or separately, also for anthropometric outcomes (BMI and bodyweight) and some cardiometabolic risk factors (systolic blood pressure, LDL cholesterol, and total cholesterol). After adjusting for small-study effects or publication bias by use of the trim-and-fill method, only the estimated change in intake of fruits, BMI, and bodyweight (eg, BMI -0.16 kg/m² [95% CI -0.29 to -0.02]) remained significant (appendix p 70).

Discussion

This systematic review and meta-analysis summarised and quantified the effect of 121 multicomponent workplace wellness programmes, based on scientific evidence accumulated in interventional trials during the past 30 years, on dietary habits, anthropometrics, and cardiometabolic risk markers. In overall pooled results, our investigation found that workplace wellness programmes influence specific dietary habits, anthropometric parameters, and cardiometabolic risk factors.

| | Country | Number of participants | Randomised | Intervention domains | Duration (months) | Effect size, cm (95% Cl) | Weight (%) |
|-------------------------------------|-------------|---------------------------|------------|-------------------------|----------------------|-----------------------------|---------------|
| Rusali (2018) ^{71*} | Malaysia | 77 | No | С | 4 🔶 | -8.90 (-13.43 to -4.37) | 0.91 |
| Morgan (2011) ⁴⁸ | Australia | 110 | Yes | ABCFGH | 3 | -5·90 (-7·60 to -4·20) | 2.72 |
| Rusali (2018) ^{71*} | Malaysia | 70 | No | BH | 4 🖛 = | -5·40 (-10·91 to 0·11) | 0.66 |
| Olafsdottir (2012) ⁶⁸ | Iceland | 61 | No | BDG | 6 | -4.80 (-7.08 to -2.52) | 2.16 |
| Raymond (2019)93 | USA | 748 | No | AFI | 60 | -4·30 (-5·72 to -2·88) | 3.02 |
| Prabhakaran (2009) ¹¹² | India | 5828 | No | BCDEI | 48 — | -3.50 (-4.30 to -2.70) | 3.65 |
| Chen (2014)73 | Taiwan | 108 | No | ACHI | 6 | -3·47 (-5·17 to -1·77) | 2.72 |
| Ryu (2017) ¹¹⁹ † | South Korea | 490 | No | BGI | 2.5 | -3·37 (-4·78 to -1·96) | 3.03 |
| Atlantis (2006)75 | Australia | 73 | Yes | ABCGI | 6 | -3·20 (-6·85 to 0·45) | 1.25 |
| Allen (2012)103 | USA | 60 | Yes | CDG | 10 | -2.79 (-5.80 to 0.21) | 1.60 |
| Limaye (2017) ⁸¹ | India | 265 | Yes | В | 12 | -2.20 (-2.88 to -1.52) | 3.76 |
| Shrivastava (2017) ⁷⁹ | India | 267 | Yes | BCG | 6 | -2.18 (-4.09 to -0.27) | 2.51 |
| Mansi (2015)100 | New Zealand | 58 | Yes | BCG | 3 | -1.90 (-4.43 to 0.63) | 1.95 |
| Lin (2017) ¹¹⁴ | Taiwan | 99 | Yes | BGI | 3 | -1.70 (-3.37 to -0.03) | 2.75 |
| Viester (2018)46 | Netherlands | 314 | Yes | ABG | 6 | -1.38 (-2.63 to -0.13) | 3.20 |
| Kwak (2010) ⁸⁴ | Netherlands | 553 | No | BCDEGI | 24 | -1.30 (-2.18 to -0.42) | 3.58 |
| Lin (2018) ⁸⁷ | Taiwan | 904 | No | AGH | 3 | -1.29 (-2.24 to -0.34) | 3.52 |
| Geaney (2016)55‡ | Ireland | 850 | No | BCE | 8 | -1.20 (-2.35 to -0.05) | 3.31 |
| Jamal (2016)57 | Malaysia | 194 | Yes | BCGHI | 6 | -1.12 (-4.00 to 1.76) | 1.68 |
| Jaime (2014) ⁹⁶ | Brazil | 720 | Yes | В | 6 | -1.05 (-1.97 to -0.13) | 3.54 |
| Balk-Møller (2017) ¹¹⁵ | Denmark | 269 | Yes | BCI | 4 | -1.05 (-2.26 to 0.16) | 3.25 |
| Geaney (2016)55‡ | Ireland | 850 | No | BCDEF | 8 | -1.00 (-4.25 to 2.25) | 1.45 |
| Ribeiro (2014)106§ | Brazil | 95 | Yes | CG | 3 | -0.87 (-1.81 to 0.07) | 3.52 |
| Cook (2001) ⁸⁸ | New Zealand | 253 | No | CDE | 6 | -0.70 (-1.64 to 0.24) | 3.52 |
| Ribeiro (2014)106§ | Brazil | 100 | Yes | ABG | 3 | , , , | 3.93 |
| Choi (2017) ¹¹⁸ | South Korea | 43 | Yes | BCG | 3 | -0.51 (-1.23 to 0.21) | 3.73 |
| Iriyama (2014)66 | Japan | 79 | Yes | BCDE | 6 | -0.50 (-2.06 to 1.06) | 2.87 |
| Viitasalo (2015)72¶ | Finland | 2312 | No | ABC | 30 | -0.12 (-1.23 to 1.00) | 3.34 |
| Viitasalo (2015) ⁷² ¶ | Finland | 2312 | No | ABC | 30 — | -0.03 (-1.17 to 1.11) | 3.32 |
| Kamioka (2009)90 | Japan | 43 | Yes | BCGI | 6 | 0.10 (-1.71 to 1.91) | 2.61 |
| Ryu (2017) ¹¹⁹ † | South Korea | 524 | No | В | 6 | 0·10 (-1·04 to 1·24) | 3.32 |
| Danguah (2017) ¹¹⁶ | Denmark | 317 | Yes | BCI | 3 - | 0.15 (-0.55 to 0.85) | 3.74 |
| Pedersen (2018) ¹¹⁷ | Norway | 202 | Yes | ABC | 5 | → 0.21 (-3.24 to 3.66) | 1.34 |
| Geaney (2016)55‡ | Ireland | 850 | No | DF | 8 | 0.50 (-3.20 to 4.20) | 1.22 |
| LaCaille (2016) ⁴⁰ | USA | 526 | No | DEGI | 12 | 0.50 (-3.29 to 4.29) | 1.18 |
| Engbers (2007) ⁸⁵ | Netherlands | 540 | No | El | 12 | 0.60 (-0.20 to 1.40) | 3.65 |
| Gerstel (2013) ⁶¹ | Switzerland | 173 | No | ACI | 12 — | 0.70 (−1.22 to 2.62) | 2.49 |
| Overall (1 ² =81.9%, p<0 | | -, 5 | | - | \sim | -1·47 (-1·96 to -0·98) | 100.00 |

Figure 3: Forest plot of waist circumference

Intervention domains correspond to screening (A), individual education (B), group education (C), food environment (D), labelling (E), financial incentives (F), physical activity (G), self-awareness (H), and others (I). Weights are from random-effects analysis. *Different intervention groups from the same study (Rusali 2018). †Different intervention groups from the same study (Ryu 2017). ‡Different intervention groups from the same study (Geaney 2016). \$Different intervention groups from the same study (Ribeiro 2014). #Different intervention groups from the same study (Viitasalo 2015).

We identified significant unexplained heterogeneity between studies, as well as potential for small-study effects or potential publication bias, that influenced the statistical significance of changes in cardiometabolic risk markers. Although magnitudes of effects were often modest, such effects are crucial and provide meaningful risk reduction when shown across populations, as opposed to through individual-focused clinical treatment.

One of the most salient findings of our comprehensive review is the wide variation in worksite settings, employee populations, intervention components, intervention durations, and outcomes assessed in studies of workplace wellness programmes. Despite comprehensive analyses to explore these factors as potential sources of heterogeneity, no definitive drivers were identified for the outcomes. One exception was a potentially larger effect on blood LDL cholesterol when interventions included changes in the food environment. A previous narrative review of 22 studies¹²² identified the food environment (eg, availability of healthy foods, cost of healthy option) as one of the intervention components most often included in workplace wellness programmes, but neither quantitative pooling of studies nor quantitative evaluation of

| | Country | Number of participants | Randomised | Intervention domains | Duration (months) | | Effect size, mg/dL (95% Cl) | Weight (%) |
|-------------------------------------|-------------|---------------------------|------------|-------------------------|----------------------|-------------------|--------------------------------|---------------|
| Allen (2012) ¹⁰³ | USA | 60 | Yes | CDG | 10 | ←── | -17·10 (-28·26 to -5·94) | 3.10 |
| Gysan (2017) ⁷⁴ | Germany | 447 | Yes | ABCI | 3.75 | • | –14·24 (–20·25 to –8·23) | 5.07 |
| Racette (2009) ⁹¹ | USA | 151 | Yes | ABCDFGI | 12 | ← | –13·95 (–17·03 to –10·87 |) 6.28 |
| Gerstel (2013)61 | Switzerland | 173 | No | ACI | 12 | — | -13·90 (-24·48 to -3·32) | 3.28 |
| Engbers (2007) ⁸⁵ | Netherlands | 540 | No | EI | 12 | ←── | –13·13 (–17·57 to –8·69) | 5.76 |
| Olafsdottir (2012) ⁶⁸ | Iceland | 61 | No | BDG | 6 | ← ■ | -11·97 (-25·67 to 1·73) | 2.42 |
| Moy (2006)77 | Malaysia | 186 | No | BCI | 24 | ← ■ → | -8·11 (-17·57 to 1·35) | 3.66 |
| Salinardi (2013)92 | USA | 133 | Yes | BC | 6 | ← ■ | -8.00 (-17.80 to 1.80) | 3.54 |
| Flannery (2012) ¹⁰⁴ | USA | 39 | Yes | BCGI | 3 | < | -7·10 (-30·98 to 16·78) |) 1.04 |
| Limaye (2017) ⁸¹ | India | 265 | Yes | В | 12 | | -4·25 (-8·08 to -0·42) | 6.01 |
| Rusali (2018)71* | Malaysia | 70 | No | BH | 4 | | -3·87 (-10·69 to 2·95) | 4.72 |
| Rusali (2018)71* | Malaysia | 77 | No | С | 4 | | -3·87 (-9·95 to 2·21) | 5.05 |
| Rowland (2018) ¹²¹ | USA | 50 | Yes | ABGH | 3 | • | -3·10 (-17·93 to 11·73) | 2.18 |
| Shrivastava (2017) ⁷⁹ | India | 267 | Yes | BCG | 6 | | -2·81 (-8·69 to 3·07) | 5.13 |
| Viitasalo (2015)72† | Finland | 2312 | No | ABC | 30 | | -1·58 (-7·57 to 4·40) | 5.08 |
| Viitasalo (2015)72† | Finland | 2312 | No | ABC | 30 | | -1·29 (-7·53 to 4·94) | 4.98 |
| Lin (2018) ⁸⁷ | Taiwan | 904 | No | AGH | 3 | | -0.76 (-3.98 to 2.46) | 6.24 |
| Lin (2017) ¹¹⁴ | Taiwan | 99 | Yes | BGI | 3 | | -0.20 (-4.82 to 4.42) | 5.68 |
| Raymond (2019)93 | USA | 748 | No | AFI | 60 | | 0.00 (-2.76 to 2.76) | 6.39 |
| Jamal (2016)57 | Malaysia | 194 | Yes | BCGHI | 6 | | 0.09 (-0.12 to 0.30) | 6.86 |
| Chen (2014)73 | Taiwan | 108 | No | ACHI | 6 | | → 1.50 (-8.13 to 11.13) | 3.60 |
| Kamioka (2009)90 | Japan | 43 | Yes | BCGI | 6 | | → 2·40 (-6·36 to 11·16) | 3.92 |
| Overall (I ² =87.5%, p<0 | 0-0001) | | | | | $\langle \rangle$ | -5·18 (-7·83 to -2·53) | 100.00 |
| | | | | | | 15 -10 -5 0 5 | 10 | |

Figure 4: Forest plot of LDL cholesterol

Intervention domains correspond to screening (A), individual education (B), group education (C), food environment (D), labelling (E), financial incentives (F), physical activity (G), self-awareness (H), and others (I). Weights are from random-effects analysis. *Different intervention groups from the same study (Rusali 2018). †Different intervention groups from the same study (Viitasalo 2015).

heterogeneity was done. We identified a stronger effect of food environment interventions on LDL cholesterol, and this observation should be viewed cautiously and might be a chance finding. Overall, our detailed review and pooled analyses support potential health benefits of workplace wellness programmes and highlight the need for greater standardisation of approaches and methods for assessing which intervention characteristics are most relevant for different populations and outcomes.

Previous meta-analyses of workplace wellness programmes have included far fewer studies. For example, Hwang and colleagues11 summarised ten studies and found significant effects on systolic blood pressure (0.66 mm Hg [95% CI 0.27 to 1.60]), diastolic blood pressure (0.63 mm Hg [0.21 to 1.06]), and BMI (0.71 kg/m² [0.15 to 1.11]), but not bodyweight or LDL cholesterol. Another systematic review and meta-analysis of workplace wellness programmes focused on obesity and included only seven studies,13 finding significant reductions in BMI with short-term interventions (<6 months: -1.26 kg/m² [95% CI -1.98 to -0.55]), but not longer-term interventions (>6 months: -1.68 kg/m² [-4.12 to 0.76]). A third meta-analysis including 24 studies¹² identified reductions in bodyweight (-2.61 kg [95% CI -3.89 to -1.33]), BMI (-0.42 kg/m² [-0.69 to -0.15]), and waist circumference (-1.92 cm [-3.25 to -0.60]), but not blood pressure, lipids, or plasma glucose; effects on dietary habits were not evaluated. By comparison with these previous studies, our investigation of 121 workplace wellness programmes builds upon and extends the previous body of evidence by retrieving and analysing a far larger number of interventional trials, providing substantially greater statistical power, generalisability of findings, and ability to assess small-study effects.

Although several significant benefits were identified in the overall pooled findings, our analysis suggested the potential for small-study effects that influence several of these outcomes. Visual inspection of funnel plots and corresponding statistical tests cannot distinguish between true differences in efficacy of smaller versus larger studies (ie, caused by other, unidentified characteristics) versus publication bias.123 Trim-and-fill methods suggested robustness of efficacy of workplace wellness programmes for fruit intake, BMI, and bodyweight, but not necessarily for other dietary factors or clinical biometrics. A sensitivity analysis restricted to randomised controlled trials showed generally similar or stronger effects for dietary and anthropometric outcomes and slightly smaller effects for cardiometabolic outcomes, although none of these differences were significantly different from the overall pooled findings. Our novel findings emphasise the need to further understand the relevance of these potential differences in additional, large, well powered intervention studies of workplace wellness programmes.

The majority of the identified studies were from highincome countries. However, non-communicable diseases are increasing rapidly in low-income and middle-income countries.¹²⁴ In line with this, we encourage the prioritisation of the development of workplace wellness programmes and their evaluation in other geographical and socioeconomic contexts. Our results highlight the generally moderate duration (about 1 year) of most workplace wellness programmes and the limited assessment of the sustainability of behaviour change after the programme ends, raising important unanswered questions about long-term effectiveness. Thus, further research with extended assessments of workplace wellness programmes should be prioritised, including assessment of the costs and cost-effectiveness of different approaches, which were often not reported in the identified studies.

This research has several strengths. This study represents the largest and most comprehensive systematic review and meta-analysis of workplace wellness programmes, allowing evaluation of various important health outcomes. We focused on multicomponent interventions, identified and recommended as most likely to be successful.^{16–18} We explored multiple potentially important sources of heterogeneity. Trimand-fill sensitivity analyses were done to quantify the potential influence of small-study effects of publication bias. The majority of studies were randomised, and findings were similar in sensitivity analyses restricted to randomised trials.

Limitations should be considered. Although study quality was assessed with a published tool used in previous meta-analyses, our quality assessment might differ from standardised methods such as GRADE or the Cochrane risk of bias tool. Study design was not identified as a source of heterogeneity in prespecified analysis, and studies were combined for the meta-analysis. This limitation should be considered when interpreting our pooled estimates. Most studies did not clearly report which outcomes were primary versus secondary in the analysis, raising the potential for biased reporting in the studies of outcomes with observed benefits. On the basis of the extensive nature of the review and number of identified studies, this research was limited to scientific literature published in English; the results might be less applicable to non-English socioeconomic settings, and potentially relevant evidence in other languages might not have been captured. As is common in meta-analyses, unexplained heterogeneity was present for most outcomes.

This study represents a characterisation of workplace wellness programmes across many settings and countries. In the past year, job dynamics have been disrupted globally by the COVID-19 pandemic, with unclear long-term repercussions. Our new findings serve as a benchmark and could be helpful in the context of considering and developing novel, more virtual workplace wellness programmes that address the shifting nature of work, especially remote work, after the pandemic. In conclusion, this systematic review and meta-analysis suggests that workplace wellness programmes can improve specific dietary, anthropometric, and cardiometabolic risk indicators, supporting their use and further investigation as effective strategies to improve cardiometabolic health.

Contributors

JLP, JS, RM, and DM conceptualised and developed the research protocol and methodology. PS, JO, JLP, and RM developed standardised data extraction tools and underlying calculation algorithms for standardised estimates. JLP, DS, EM, IU, JS, JHYW, EB, and JO reviewed the literature, selected eligible studies, and performed data extraction. JLP and DS did the statistical analysis. JLP, DS, and DM interpreted results, and drafted, reviewed, and edited the manuscript. JLP and DM supervised the study. JLP and DS accessed and verified the data. All authors have access to the underlying data. JLP and DM were responsible for the decision to submit the manuscript for publication. All authors have revised and approved the final text. JLP, as guarantor of this manuscript, affirms that it is an honest, accurate, and transparent account of the study being reported; that no important aspects of the study have been omitted; and that any discrepancies from the study as planned have been explained and registered.

Declaration of interests

JS is a current employee of General Mills, a food manufacturer; her contribution to this work took place during her affiliation with Tufts University. Her salary at the time of this work was supported by a postdoctoral fellowship from the Canadian Institutes of Health Research. RM reports grants to her institution from the Bill & Melinda Gates Foundation, Nestle, and Danone, outside the submitted work; and consulting fees from Development Initiatives, outside the submitted work. All other authors declare no competing interests.

Data sharing

The data used in this study may be accessible under request by contacting the corresponding author.

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